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SOME ASYMPTOTIC RESULTS FOR OCCUPANCY PROBLEMS

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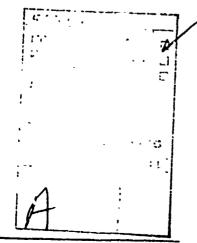
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ABSTRACT

Consider a situation in which balls are falling into N cells with arbitrary probabilities. Limit distributions for the number of empty cells are considered when $N+\infty$ and the number of balls $n+\infty$ so that $n/N+\infty$. Limit distributions for the number of balls to achieve exactly b empty cells are obtained when $N+\infty$ for b fixed or $b+\infty$ so that b/N+0.

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SOME ASYMPTOTIC RESULTS FOR OCCUPANCY PROBLEMS

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1. Introduction.

Suppose that balls are thrown independently of each other into N cells, so that each ball has the probability p_k of falling into the kth cell, $p_1 + \ldots + p_N = 1$. Let Y_n denote the number of empty cells after n throws and let T_b denote the throw on which for the first time exactly b cells remain empty, $0 \le b < N$. The symmetrical case $p_1 = \ldots = p_N = 1/N$ is discussed in e.g. Felier (1968), see occupancy or waiting time problems.

Depending on how b, n, $N \rightarrow \infty$, different asymptotic distributions for Y_n and T_b can be obtained, see e.g. Holst (1971) and for the symmetric case see e.g. Samuel-Cahn (1974). In this paper some remaining problems are investigated for the nonsymmetrical case.

To give precise meanings of the limits obtained, double sequences e.g. $(p_{kN})_N, (Y_{nN})_N$ are considered. But in order to simplify the notation the extra index N will usually be omitted.

2. A bounded number of empty cells.

The following limit theorem for Y_n , the number of empty cells after n throws, was proved by Sevastyanov (1972).

Theorem 1. If the p's are such that

(2.1)
$$\max_{1 \le k \le N} (1 - p_k)^n \rightarrow 0$$

and

(2.2)
$$E(Y_n) = \sum_{k=1}^{N} (1 - p_k)^n \rightarrow m < \infty ,$$

then

(2.3)
$$P(Y_n = y) \rightarrow m^Y \cdot e^{-m}/y!$$
,

or equivalently

(2.4)
$$Y_n \Rightarrow Po(m)$$
, when $N \to \infty$.

Remark. When the p's are equal an expression for $P(Y_n = y)$ can be obtained from which (2.3) can be derived by elementary methods, see e.g.

Feller (1968). In this case (2.1) and (2.2) are replaced by

$$(2.5) N \cdot \exp(-n/N) \to m < \infty$$

or

$$(2.6) n/N - \log N \rightarrow -\log m > -\infty.$$

For $\mathbf{T}_{\mathbf{b}}$, the number of balls until \mathbf{b} empty cells remain, the limit distribution is given by:

Theorem 2. If b is a fixed integer and for some fixed numbers C and D,

(2.7)
$$0 < C \le Np_k \le D < \infty$$
, for all k and N,

then, when $N \rightarrow \infty$,

(2.8)
$$\sum_{k=1}^{N} (1 - p_k)^{T_b} \implies \frac{1}{2} \chi^2(2(b+1)),$$

and

(2.9)
$$\sum_{k=1}^{N} \exp(-T_b p_k) \Rightarrow \frac{1}{2} \chi^2(2(b+1)).$$

Before proving the theorem the following functions are considered:

(2.10)
$$f(t) = f_{N}(t) = \sum_{k=1}^{N} (1-p_{k})^{t}, \quad t > 0,$$

and

(2.11)
$$g(t) = g_N(t) = \sum_{k=1}^{N} \exp(-tp_k).$$

<u>Lemma 1.</u> If Condition (2.7) is satisfied, y > 0 is a fixed number, and $t = t_N = t(y)$ is defined by the equation

$$(2.12)$$
 $f(t) = y$,

then

$$(2.13) 0 < C \le \lim_{N \to \infty} \inf N \log N/t_N \le \lim_{N \to \infty} N \log N/t_N \le D < \infty$$

and when $N \rightarrow \infty$

$$(2.14) f([t]) \rightarrow y,$$

(2.15)
$$\max_{1 \le k \le N} (1 - p_k)^{[t]} \to 0 ,$$

(2.16)
$$g(t)$$
 and $g([t]) \rightarrow y$.

where [t] denotes the integer part of t.

<u>Lemma 2.</u> If f is replaced by g and g by f in Lemma 1, then the same conclusions hold.

Proof of Lemma 1. From Condition (2.7), it follows that

(2.17)
$$y = \sum_{k=1}^{N} (1 - p_k)^{t} \ge N \cdot (1 - D/N)^{t}.$$

Hence for $\varepsilon > 0$ and N sufficiently large

(2.18)
$$\log y \ge \log N - t \cdot (D+\epsilon)/N$$

and therefore

(2.19) D+ ϵ = (D+ ϵ) lim. (1/(1-log y/logN)) \geq lim sup N log N/t N $\rightarrow \infty$ which proves the right inequality of (2.13).

To prove the left inequality of (2.13) the following estimate follows from (2.7):

(2.20)
$$y = \sum_{1}^{N} (1-p_k)^t \leq N \cdot (1-C/N)^t$$
,

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or

(2.21)
$$\log y \leq \log N - t \log(1 - C/N) \leq \log N - tC/N.$$

From this it follows that

(2.22)
$$C = C \lim_{N \to \infty} (1 - \log y/\log N)^{-1} \le \lim_{N \to \infty} \inf_{N \to \infty} N \log N/t_{N}.$$

To prove (2.14) we observe that

$$(2.23) (1 - p_k)^{t-1} \ge (1 - p_k)^{[t]} \ge (1 - p_k)^t,$$

and using (2.7)

$$(2.24) \quad (1 - D/N)^{-1} \sum_{1}^{N} (1-p_k)^{t} \geq \sum_{1}^{N} (1-p_k)^{[t]} \geq \sum_{1}^{N} (1-p_k)^{t} ,$$

or from (2.12)

$$(2.25) (1 - D/N)^{-1} y \ge f([t]) \ge y.$$

From which (2.14) follows.

Combining (2.7) and (2.13) give for some $\frac{K}{1} > 0$ and N sufficiently large that

(2.26)
$$\max(1-p_k)^{[t]} \le (1-C/N)^{[t]} \le (1-C/N)^{[t]} \to 0, N \to \infty$$
 which proves (2.15).

Using (2.7) and (2.13) it follows that for some constant K

(2.27)
$$|1 - e^{-tp} k/(1 - p_k)^t| \le K \cdot \log N/N$$
,

and therefore

(2.28)
$$|f(t) - g(t)| \leq \sum_{1}^{N} (1 - p_k)^t \cdot |1 - e^{-tp_k} / (1 - p_k)^t |$$

$$\leq K \sum_{1}^{N} (1 - p_k)^t \log N / N = K y \log N / N \rightarrow 0 ,$$

which proves (2.16).

Proof of Lemma 2. The proof is essentially the same as that for Lemma 1.

Proof of Theorem 2. From the definitions it follows that

(2.29)
$$Y_n \le b \iff T_b \le n$$
,

and therefore

(2.30)
$$P(Y_n \le b) = P(T_b \le n) = P(f(T_b) \ge f(n))$$
.

Let y > 0 be fixed and define n = [t] with t = t(y) as in Lemma 1. According to Lemma 1 the assumptions of Theorem 1 are satisfied. Hence

(2.31)
$$P(f(T_b) \ge y) = P(Y_n \le b) \rightarrow P(Y \le b),$$

where Y is Pc(y). Furthermore it is well-known that

(2.32)
$$P(Y \le b) = P(\frac{1}{2}\chi^2(2(b+1)) \ge y)$$
.

(2.31) and (2.32) prove (2.8). Using Lemma 2, the assertion (2.9) follows.

Remark. When the p's are equal the theorem can be written

(2.33)
$$\sqrt{1 \cdot (1 - 1/N)}^{T_b} = \frac{1}{2} \chi^2 (2(b+1))$$
,

and therefore

(2.34)
$$T_b/N - \log N \Rightarrow \log(\frac{1}{2}\chi^2(2(b+1)))$$
.

This result was found by Baum and Billingsley (1965) using complicated calculations. Using the result in Feller (1968) and the method of proof of Theorem 2, (2.33) and (2.34) follows. A consequence of (2.34) is

(2.35)
$$T_b/N \log N \rightarrow 1$$
, in probability, as $N \rightarrow \infty$.

Now (2, 35) will be generalized. First introduce the distribution function

(2.36)
$$H_N(x) = \# (p_k : Np_k \le x)/N$$
.

<u>Lemma 3.</u> If $t = t_N = t(y)$ is defined by

(2.37)
$$g(t) = g_{N}(t_{N}) = y > 0$$
,

and there exists a distribution function H(x) on [C,D] such that

(2.38)
$$H_N(x) \rightarrow H(x)$$
, $N \rightarrow \infty$,

and

$$(2.39) 0 < C = \inf\{x : H(x) > 0\},\$$

then for $1/C > \epsilon > 0$, when $N \rightarrow \infty$,

(2.40)
$$g_N((\epsilon + 1/C)(N \log N)) \to 0$$
,

and

(2.41)
$$g_{N}(-\varepsilon + 1/C)(N \log N)) \rightarrow +\infty.$$

Proof. From the definitions it follows that

(2.42)
$$0 < y = g_{N}(t_{N}) = N \cdot \int_{C}^{D} \exp(-t_{N}x/N) dH_{N}(x) = \int_{C}^{D} \exp((1-t_{N}x/N) \log N) \log N dH_{N}(x).$$

Consider

$$(2.43) \ g_N((\epsilon+1/C) \ N \log N) = \int_C^D \exp((1-x(1+\epsilon C)/C) \log N) dH_N(x) \ .$$
 Now for $C \le x \le D$ it is true that $1 - x(1+\epsilon C)/C < 0$ and therefore the exponent in (2.43) is negative so the integral tend to 0 when $N \to \infty$, which proves (2.40).

For proving (2.41) consider

(2.44)
$$g_N((-\epsilon + 1/C) \ N \log N) = \int_C^D \exp((i - x(1 - \epsilon C)/C) \log N) dH_N(x)$$
.
For $C \le x \le C/(1-C\epsilon)$ the exponent is positive and as the integrand is positive (2.44) could be estimated by

(2.45)
$$\int_{C}^{C/(1-C\varepsilon)} \exp((1-x(1-\varepsilon C)/C)\log N) dH_{N}(x) \to +\infty$$
 by Condition (2.39).

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Corollary to Theorem 2. If the Conditions (2.38) and (2.39) are satisfied then

(2.46)
$$T_b/N \log N \rightarrow 1/C$$
, in probability, $N \rightarrow \infty$.

<u>Proof.</u> Let $\epsilon_1 > 0$ and $\epsilon_2 > 0$ be given. Take a $\delta > 0$ so that

(2.47)
$$P(\frac{1}{2}\chi^2(2(b+1)) < \delta) < \epsilon_2/2$$
.

For N sufficiently large it follows from Theorem 2 that

(2.48)
$$P(g_{N}(T_{b}) < \delta) < \epsilon_{2}/2$$

and from Lemma 3 that

(2.49)
$$g_N((\epsilon_1 + 1/C)(N \log N)) < \delta$$
.

Hence

(2.50)
$$P(T_b/N \log N > \epsilon_1 + 1/C) =$$

$$P(g_N(T_b) < g_N((\epsilon_1 + 1/C)(N \log N)) \le 1 + 1/C$$

$$P(g_N(T_b) < \delta) < \epsilon_2/2 .$$

In a similar way it is proven that

(2.51)
$$P(T_b/N \log N < -\epsilon_1 + 1/C) < \epsilon_2/2$$
.

Hence for N sufficiently large

(2.52)
$$P(|T_b/N \log N - 1/C| > \epsilon_1) < \epsilon_2$$
.

Thus the assertion is proved.

3. A small fraction of empty cells.

As above, Y_n denotes the number of empty cells after n throws.

Theorem 3. 1f

$$(3.1) 0 < C \le Np_k \le D < \infty, \text{ for all } k \text{ and } N,$$

$$(3.2) n/N \to \infty,$$

and

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(3.3)
$$f(n) = E(Y_n) = \sum_{k=1}^{N} (1 - p_k)^n \to +\infty,$$

then, when $n \rightarrow \infty$,

$$(3.4) \qquad (Y_n - f(n))/(f(n))^{\frac{1}{2}} \implies N(0,1) ,$$

and

(3.5)
$$(Y_n - g(n))/(g(n))^{\frac{1}{2}} => N(0,1),$$

where

(3.6)
$$g(n) = \sum_{k=1}^{N} \exp(-np_k)$$
.

Proof. Using (3.1) and (3.3) it follows that

(3.7)
$$\sum_{1}^{N} (1 - p_{k})^{n} \leq N \cdot (1 - C/N)^{n} \rightarrow +\infty,$$

hence

(3.8)
$$n/N \log N = O(1)$$
.

Using (3.1), (3.2), and (3.8) give

$$|f(n) - g(n)| \leq \sum_{l}^{N} \exp(-np_{k}).$$

$$\cdot |\exp(n \log (l - p_{k}) + np_{k}) - 1| \leq$$

$$\leq \sum_{l}^{N} \exp(-np_{k}) \cdot K \cdot n/N^{2} \leq$$

$$\leq K \cdot (n/N) \cdot \exp(-C n/N) \rightarrow 0.$$

Hence it is sufficient to prove (3.5). This will be established using convergence of characteristic functions.

In Holst (1971) p. 1672 the characteristic function of Y_n is given by

(3.10)
$$E(\exp(itY_n)) = (n!/2\pi i N^n) \cdot \frac{1}{|x|} (1 + (e^{it}-1)\exp(-Np_k z)) dz$$

$$|z| = \pi/N$$

Using Stirling's formula and changing to polar coordinates it follows that

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(3.11)
$$E(\exp(it(Y_{n} - \mu)/\sigma)) = (1 + o(1)).$$

$$\cdot \int_{-\pi}^{\pi} (n/2\pi)^{\frac{1}{2}} \cdot \exp(n(e^{i\theta} - 1 - i\theta)).$$

$$\cdot \frac{N}{||} (\exp(-it e^{-np} k/\sigma) \cdot (1 + (e^{it/\sigma} - 1)\exp(-np_{k} e^{i\theta})))d\theta$$

$$= (1 + o(1)) \cdot \int_{-\pi}^{\pi} h_{n}(\theta, t)d\theta ,$$

where

(3.12)
$$\mu = \sigma^2 = g(n) = \sum_{k=1}^{N} \exp(-np_k), \quad \sigma > 0.$$

The integral will be studied by the same method as in Holst (1971).

Take 0 < a < 1/6 and split the interval $-\pi \le \theta \le \pi$ into

(3.13)
$$A = \{\theta ; a \leq |\theta| \leq \pi \},$$

(3.14)
$$B = \{ \theta; n^{a-\frac{1}{2}} \le |\theta| < a \},$$

and

(3.15)
$$C = \{\theta ; |\theta| < n^{a-\frac{1}{2}}\}.$$

From Lemmas 4-6 below it follows that

(3.16)
$$E(\exp(it(Y_n - \mu)/\sigma) = (1 + o(1)).$$

$$(\int_A h_n + \int_B h_n + \int_C h_n) \to 0 + 0 + \exp(-t^2/2), \quad n \to \infty.$$

By the continuity theorem for characteristic functions assertion (3.5) is proved, and thus the theorem.

With the same conditions as in Theorem 3 the following lemmas hold.

Lemma 4. For every fixed real number t

(3.17)
$$\int_{A} h_{n}(\theta,t)d\theta \rightarrow 0 , \quad n \rightarrow \infty .$$

Proof. As $n/N \to \infty$ and $\sigma \to \infty$ it follows that

$$|\int_{A} | \leq K_{1} \cdot n^{\frac{1}{2}} e^{-n} \cdot \int_{A} \prod_{l} |\exp(np_{k}e^{i\theta}) + e^{it/\sigma} - l| d\theta$$

$$\leq K_{2} n^{\frac{1}{2}} e^{-n} \prod_{l} (\exp(np_{k} \cos a) + o(l))$$

$$\leq K_{2} n^{\frac{1}{2}} e^{-n} 2^{N} e^{n\cos a} \rightarrow 0.$$

Lemma 5 For ever liked real number t

(3.19)
$$\int_{B} \dot{n}_{11}(\cdot,t)dA \rightarrow 0 , \quad n \rightarrow \infty .$$

Proof. From the assumptions, it follows that there exist positive numbers

 $K_3 - K_9$ such that

$$|\int_{B} | \leq K_{3} n^{\frac{1}{2}} e^{-n} \int_{B} \frac{N}{||} (\exp(np_{k} \cos \theta) + O(1/\sigma)) d\theta$$

$$\leq K_{4} n^{\frac{1}{2}} e^{-n} \frac{N}{||} \exp(np_{k} \cos n^{a-\frac{1}{2}}) \cdot$$

$$\cdot (1 + K_{5} \cdot \exp(-K_{6} n/N)/\sigma)$$

$$\leq K_{7} n^{\frac{1}{2}} e^{-n} \exp(n(1 - K_{8} n^{2a-1}))$$

$$\leq \exp(-K_{6} n^{2a}) \rightarrow 0, \quad n \rightarrow \infty.$$

<u>Lemma 6.</u> For every fixed real number t,

(3.21)
$$\int_{C} h_{n}(\theta,t)d\theta \rightarrow \exp(-t^{2}/2), \quad n \rightarrow \infty.$$

Proof. Expanding in series gives

(3.22)
$$\log h_n(\theta,t) = -n \theta^2/2 + o(1)$$

$$+ \sum_{l}^{N} (\log (l + \exp(-np_k e^{i\theta}) (e^{it/\sigma} - l)) - it \exp(-np_k)/\sigma) + \frac{1}{2} \log(n/2\pi).$$

Now, when $n \rightarrow \infty$,

(3.23)
$$\sum_{1}^{N} |\exp(-2np_{k} e^{i\theta})(e^{it/\sigma} - 1)^{2}|$$

$$= o(1) \cdot \sum_{1}^{N} \exp(-np_{k})/\sigma^{2} = o(1),$$

and therefore

(3.24)
$$\sum_{1}^{N} (\log (1 + ...) - ...)$$

$$= \sum_{1}^{N} (\exp(-np_{k} e^{i\theta})(e^{it/\sigma} - 1) - it \exp(-np_{k})/\sigma) + o(1).$$

Furthermore, using (3.8), (3.9) and the assumptions; it follows that

(3.25)
$$\sum_{1}^{N} \exp(-np_{k} e^{i\theta})/\sigma^{2} \to 1,$$

and therefore (3.24) can be written

(3.26)
$$\sum_{1}^{N} (...) = \sum_{1}^{N} (\exp(-np_{k} e^{i\theta}))(it/\sigma - t^{2}/2\sigma^{2})$$

$$- it \exp(-np_{k})/\sigma) + o(1)$$

$$= it \sum_{1}^{N} (\exp(-np_{k}(e^{i\theta} - 1)) - 1) \exp(-np_{k})/\sigma$$

$$- t^{2}/2 + o(1).$$

Now, when $n \rightarrow \infty$

(3.27)
$$\sum_{1}^{N} (np_{k})^{2} \theta^{2} \exp(-np_{k})/\sigma \leq$$

$$\leq K_{1} (n/N)^{2} n^{2a-1} N^{\frac{1}{2}} \exp(-K_{2} n/N) \rightarrow 0.$$

From this it follows that

(3.28)
$$\sum_{1}^{N} (...) = \theta t \sum_{1}^{N} n p_{k} \exp(-np_{k})/\sigma - t^{2}/2 + o(1).$$

Hence for θ in C,

(3.29)
$$\log h_{n}(\theta, t) - \frac{1}{2} \log(2\pi/n) = -n\theta^{2}/2 + \theta t \sum_{l}^{N} n p_{k} \exp(-n p_{k})/\sigma$$

$$- t^{2}/2 + o(1) = -(n^{\frac{1}{2}}\theta - t \sum_{l}^{N} n^{\frac{1}{2}} p_{k} \exp(-n p_{k})/\sigma)^{2}/2$$

$$- t^{2}(1 - (\sum_{l}^{N} n^{\frac{1}{2}} p_{k} \exp(-n p_{k})/\sigma)^{2})/2 + o(1) .$$

Now, when $n \to \infty$,

(3.30)
$$\sum_{1}^{N} n^{\frac{1}{2}} p_{k} \exp(-n p_{k}) / \sigma \leq K_{3} n^{\frac{1}{2}} N^{-1} \cdot N^{\frac{1}{2}}$$

$$\cdot \exp(-K_{4} n / N) \rightarrow 0.$$

Thus with $\psi = n^{\frac{1}{2}\theta}$ the integral (3.21) can be written

(3.31)
$$\int_{C} h_{n} = \int_{|\psi| \le n^{a}} (2\pi)^{-\frac{1}{2}}$$

$$\cdot \exp(-(\psi - o(1))^{2}/2 - t^{2}/2 + o(1)) d\psi,$$

which converges to $\exp(-t^2/2)$ when $n \to \infty$.

4. The waiting time for a small fraction.

As above let T_b denote the number of balls thrown until exactly $b = b_N$ cells remain empty. Let t_b be the unique solution of the equation $(4.1) \qquad b = g(t_b) = \sum_{k=1}^{N} \exp(-t_b p_k).$

Theorem 4. If, when $N \to \infty$,

$$(4.2) b_N \to +\infty,$$

$$(4.3) b_N/N \rightarrow 0,$$

and

$$(4.4) C < C \le NP_k \le D < \infty, \text{ for all } k \text{ and } N,$$

then

(4.5)
$$b_{N}^{-\frac{1}{2}}(T_{b} - t_{b}) \sum_{k=1}^{N} p_{k} \exp(-t_{b}p_{k}) \Rightarrow N(0,1).$$

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<u>Proof.</u> From the assumptions it follows that

(4.6)
$$C b/N \leq \Delta = \sum_{k=1}^{N} p_{k} \exp(-t_{b}p_{k}) \leq D b/N.$$

Thus for N sufficiently large

$$(4.7) 0 < C \leq \Delta \cdot N/b \leq D < \infty.$$

As in the proof of Theorem 2 the following relation holds

(4.8)
$$P((T_b - t_b) \Delta/b^{\frac{1}{2}} \le x) = P(Y_n \le b)$$
,

where

(4.9)
$$n = [t_b + x b^{\frac{1}{2}}/\Delta].$$

It is seen that

(4.10)
$$g(n) (l + o(1)) = g(t_b + x b^{\frac{1}{2}}/\Delta)$$

$$= \sum_{k} \exp(-t_b p_k) \cdot (l - x p_k b^{\frac{1}{2}}/\Delta + O(l/L_i)$$

$$= b - x \cdot b^{\frac{1}{2}} + O(l),$$

and thus

$$(4.11) g(n) \rightarrow +\infty,$$

and from (3.9) it follows that

$$(4.12) f(n) \rightarrow +\infty.$$

Furthermore,

(4.13)
$$b = g(t_b) \ge N \exp(-D t_b/N),$$

implying that

$$(4.14) t_h /N \rightarrow +\infty ,$$

and therefore

$$(4.15) n/N \to +\infty.$$

Hence the assumptions of Theorem 3 are fulfilled and (4.8) and (4.10) give

$$(4.16) P(T_b - t_b) \Delta / b^{\frac{1}{2}} \le x) = P(Y_n \le b) =$$

$$= \Phi ((b - g(n)) / (g(n))^{\frac{1}{2}}) + o(1) =$$

$$= \Phi ((x b^{\frac{1}{2}} + O(1)) / (b(1 + o(1)))^{\frac{1}{2}}) + o(1) \rightarrow \Phi(x),$$

where $\Phi(x)$ is the standardized normal distribution function. This proves the theorem.

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